

Interpretation of a Commercial Bovine Paratuberculosis Enzyme-Linked Immunosorbent Assay by Using Likelihood Ratios

M. T. Collins*

School of Veterinary Medicine, University of Wisconsin, Madison, Wisconsin 53706

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Evidence-based medicine encourages the use of quantitative diagnostic test results to estimate the probability of a particular diagnosis. Likelihood ratios (LRs) are among the best tools for maximizing the diagnostic information gained from diagnostic assays that provide results on a continuous scale. They provide the odds that an animal with a particular test result actually has the disease in question based on the magnitude of the test result. A commercial enzyme-linked immunosorbent assay (ELISA) was used to test sera from 143 dairy cattle infected with *Mycobacterium paratuberculosis* and 2,974 cattle free of this infection. This assay transforms ELISA reader optical density values into sample-to-positive (S/P) ratios. The LR was calculated for S/P results from 0.00 to 1.00 at 0.05-S/P unit intervals. LRs were directly but not linearly correlated with ELISA S/P ratios ($r^2 = 0.94$). The mathematical function describing the relationship between the ELISA S/P ratio and the LR was $LR = 265 \times (S/P \text{ value})^{2.03}$. LRs were also directly related to the frequency of animals testing positive for paratuberculosis by fecal culture and other serologic tests. Based on these LRs, guidelines for interpretation and application of this ELISA for the diagnosis and control of paratuberculosis in dairy cattle herds are recommended.

The principal goal of a diagnostic test is to help practitioners increase the probability of a correct diagnosis. Predictive values are useful in this regard but require the estimation of disease prevalence (i.e., pre-test probability of disease) in the population (27) and only utilize categorical test results, positive or negative. Food-producing animals exist in numerous discrete populations (herds or flocks), and the prevalence of disease can differ greatly among them. Infection prevalence significantly affects predictive values, i.e., the positive predictive value of tests is low when prevalence is low, and the negative predictive value is low when disease prevalence is high, thus the predictive values of tests can vary significantly among herds.

Evidence-based medicine (EBM) is the scrupulous, explicit, and judicious use of the best evidence available in making decisions about the care of individual patients. The practice of EBM means integrating clinical expertise with the best available clinical evidence from systematic diagnostic research (19). Additional information on EBM is available at the Learning and Information Services website on EBM (<http://www.herts.ac.uk/lis/subjects/health/ebm.htm>). The likelihood ratio (LR) is one example of external clinical evidence and a powerful tool in EBM. LRs give the same information as predictive values but can be used independent of pre-test disease prevalence (5).

The purpose of the study was to generate an algorithm for the calculation of LRs for subclinical *Mycobacterium paratuberculosis* infection in dairy cattle (John's disease) based on an enzyme-linked immunosorbent assay (ELISA) for serum

antibodies and to show its utility in the control of this economically important and potentially zoonotic infectious disease (1, 6, 10, 17).

MATERIALS AND METHODS

Serum samples. Samples originated from two well-defined dairy cattle populations. Sera from 143 subclinically *M. paratuberculosis*-infected cows were part of a previously described specimen repository (23). The case definition for *M. paratuberculosis*-infected cattle was the isolation of *M. paratuberculosis* by fecal culture and/or histopathologic evidence of infection. Sera from uninfected cows included 760 samples from U.S. dairy cattle and 2,214 samples from Dutch dairy cattle. These cattle were from herds free of *M. paratuberculosis* infection as defined by a minimum of three negative annual whole-herd (all cattle, ≥ 2 years old) fecal cultures.

ELISA. An *M. paratuberculosis* antibody test kit (IDEXX Laboratories, Inc., Westbrook, Maine) was used to test all 3,117 sera according to the manufacturer's instructions. With this kit, optical density (OD) values were transformed to S/P ratios based on the OD for the serum sample together with those for the negative and positive controls provided with the kit by using the following equation: $S/P \text{ ratio} = (OD \text{ of sample} - OD \text{ of negative control}) / (OD \text{ of positive control} - OD \text{ of negative control})$. All assays were run in duplicate. Any assay with a between-well coefficient of variation of $>10\%$ was repeated, and the second result was used for data analysis.

Other tests for paratuberculosis. ELISA results were compared to those of other tests for paratuberculosis run on samples collected at the same time from the same cattle. Fecal culture was done both by the BACTEC system (2) and by using conventional solid medium (Herrold's egg yolk agar) (29). Serum antibody measurements were done by the complement fixation test (12), agar gel immunodiffusion assay (20), and another commercial ELISA kit (Paracheck; Biocor Animal Health, Omaha, Nebr.) (3). These tests were performed independently and simultaneously at the time the original samples were collected. Results of these other diagnostic tests for paratuberculosis were reported previously and not done specifically for the present study (22–24).

Data analysis. Frequency distributions for S/P values on the sera from infected and noninfected populations were tabulated in intervals of 0.05 S/P units. At each interval, the sensitivity (Se), specificity (Sp), and LR $[Se/(1-Sp)]$ of the ELISA were calculated. Linear and nonlinear regression analysis was used to determine the equation describing the line best fitting the plot of the S/P cutoff value versus LR (Lotus Freelance Graphics release 9.6 for Windows; Lotus

* Mailing address: School of Veterinary Medicine, University of Wisconsin—Madison, 2015 Linden Dr., Madison, WI 53706-1102. Phone: (608) 262-8457. Fax: (608) 265-6463. E-mail: mcollin5@facstaff.wisc.edu.

TABLE 1. *M. paratuberculosis* serum antibody ELISA Se, Sp, and LR values by ELISA S/P value

ELISA S/P cutoff value	No. of sera above cutoff from:		% Sp ^c	% Se ^d	LR ^e
	<i>M. paratuberculosis</i> -free cattle ^a	<i>M. paratuberculosis</i> -infected cattle ^b			
0.00	1,904 ^f	116 ^f	36.0	88.6	1
0.05	995	94	66.5	71.8	2
0.10	545	77	81.7	58.8	3
0.15	331	71	88.9	54.2	5
0.20	198	70	93.3	53.4	8
0.25	91	63	96.9	48.1	16
0.30	71	59	97.6	45.0	19
0.35	54	57	98.2	43.5	24
0.40	46	55	98.5	42.0	27
0.45	36	51	98.8	38.9	32
0.50	27	51	99.1	38.9	43
0.55	21	50	99.3	38.2	54
0.60	19	50	99.4	38.2	60
0.65	15	47	99.5	35.9	71
0.70	14	46	99.5	35.1	75
0.75	10	46	99.7	35.1	104
0.80	8	44	99.7	33.6	125
0.85	6	44	99.8	33.6	166
0.90	5	42	99.8	32.1	191
0.95	5	42	99.8	32.1	191
≥1.00	3	41	99.9	31.3	310

^a $n = 2,974$.

^b $n = 143$.

^c Number below cutoff/total number of *M. paratuberculosis*-free cattle tested.

^d Number above cutoff/total number of *M. paratuberculosis*-infected cattle tested.

^e LR, the odds that a cow with this ELISA S/P value is *M. paratuberculosis* infected were determined using the equation $LR = Se/(1 - Sp)$.

^f Sera from some tested cattle produced OD values below that of the negative control, thus giving S/P values less than zero; hence, this number did not equal the total number of sera tested in this population.

Development Corp.). The regression model providing the highest r^2 value was considered to best fit the data.

RESULTS

There was a small difference in ELISA false-positive rates between the U.S. and Dutch cattle free of *M. paratuberculosis*; maximum ELISA S/P values were 1.15 and 1.10 for the two groups, respectively, and the Sp values at the manufacturer's recommended S/P cutoff values of 0.25 were 98.7 and 96.3% for the U.S. and Dutch *M. paratuberculosis*-free cattle, respectively. Pooling the data for these two populations improved the precision of LR estimates and potentially made the findings more universally applicable across countries.

Estimated Se, Sp, and LR increased with increasing ELISA S/P cutoff values (Table 1). Regression analysis of ELISA S/P ratio category versus LR was attempted using linear, exponential, logarithmic, and power functions. The S/P-to-LR relationship that best fit the data was described by the power function $LR = 265 \times (S/P \text{ value})^{2.03}$. This function fitted the data with an r^2 value of 0.94 (Fig. 1). Other regression functions had far lower r^2 values.

When the 143 *M. paratuberculosis*-infected cattle were arbitrarily clustered into five groups according to ELISA S/P levels (levels found useful based on clinical experience), a relationship between the magnitude of ELISA S/P value and the rate

Likelihood ratio

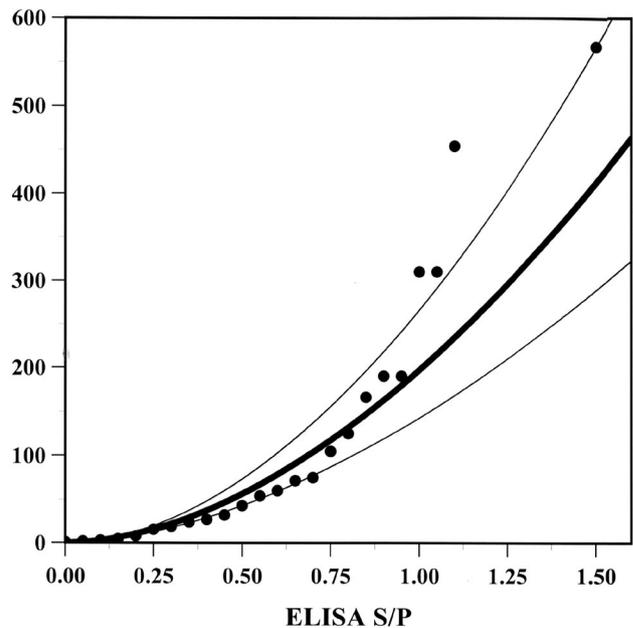


FIG. 1. Nonlinear regression of ELISA S/P values and LR (bold line) with 95% confidence intervals (thin lines).

at which cattle tested positive on other tests for paratuberculosis in historical data was apparent (Table 2).

DISCUSSION

Control of paratuberculosis in dairy cattle herds requires hygienic measures to limit opportunities for transmission of the infection from cows to calves in combination with the management of cattle most likely to be infectious (7, 8, 28). Infected cattle should not provide colostrum or milk to calves, and their manure should not be allowed to contaminate feed, water, or the environment. This is particularly true for the pens in which calves are born and the location on the farm where calves are raised. In addition, as many of the infected cows as possible should be culled from the herd when economically feasible. Because the majority of *M. paratuberculosis*-infected cattle are infectious (shedding the organism in their manure, colostrum, and milk) but clinically normal, laboratory diagnostics are needed to identify them.

Culture of feces to diagnose *M. paratuberculosis* infection with conventional culture media, such as Herrold's egg yolk agar, requires 8 to 16 weeks (29). Laboratories typically charge \$12 to 25 per sample. A liquid-culture-based detection system such as the Trek ESP system and the BACTEC system are able to shorten the detection time to 4 to 8 weeks but are not less costly than conventional culture when the costs of isolate identification are considered (2, 9). Genetic *M. paratuberculosis* detection technology coupled with PCR methods theoretically should enhance detection Se and considerably shorten the time to detection. However, commercial tests have yet to attain the analytical Se of culture methods, are roughly twice as expensive, and are difficult to scale up for handling large sample numbers (30).

TABLE 2. Percentage of positive results by other diagnostic methods when *M. paratuberculosis*-infected cattle were grouped by ELISA S/P range

IDEXX ELISA S/P value	No. of cattle tested	% Positive results in fecal culture by:		% Serum antibody by:		
		Modified BACTEC system ^a	Conventional methods on HEY ^b	AGID ^c	CF ^d	Alternative commercial antibody ELISA ^e
<0.10	65	26	20	0	5	5
0.10–0.24	13	39	46	15	23	31
0.25–0.39	9	33 ^f	22	22	44	44
0.40–0.99	15	67	33	20	7	67
≥1.00	41	93	80	76	68	95

^a Modified BACTEC 12B media with filter-concentrated fecal specimens (2).

^b Decontamination with hexadecyl cetylpyridinium chloride, concentration by sedimentation, and inoculation of Herrold's egg yolk agar (HEY) as performed by the Wisconsin Veterinary Diagnostic Laboratory, Madison, Wis. (29).

^c Agar gel immunodiffusion assay (AGID) done by Rapid Johne's test (ImmuCell Corp., Portland, Maine) (20, 21).

^d Complement fixation test (CF) performed by the Wisconsin Veterinary Diagnostic Laboratory with reagents supplied by the National Veterinary Services Laboratory, Ames, Iowa (12).

^e Paracheck (3).

^f Confidence intervals for the reported percentages were large, i.e., up to ±16% for ELISA categories with very few cattle.

Serology provides a cost-effective alternative to organism detection-based diagnostic methods for bovine paratuberculosis. ELISA-based methods have the highest Se of serologic tests for paratuberculosis (24), plus they offer the kind of low-cost and high-throughput process (>1,000/day) needed to serve the dairy industry. A disadvantage of ELISAs for paratuberculosis is that assay Sp is less than that for fecal culture (considered to be 100%) (13, 25, 31) and the economic consequences of mistakenly culling a cow due to false-positive test results are high (roughly \$1,300/cow based on the average price of replacement cattle in the United States in 2001 and the average salvage value of culled dairy cow).

Traditional ELISA interpretation is dichotomous (positive or negative) based on a single-assay cutoff value designed to optimize assay Se and Sp. Use of multilevel LRs capitalizes on the assay's ability to report results on a continuous scale, thereby enhancing the amount of diagnostic information gained. LRs quantify the probability of an accurate diagnosis. Diagnostic probabilities generated from LRs, with or without use of pre-test probabilities, can be integrated with the economic impact of actions taken based on test results such as culling. Thurmond et al. nicely demonstrated this by an ELISA for *Neospora caninum* infection in dairy cattle (26).

LRs are derived from Se and Sp estimates. These traditional measures of test accuracy for infectious diseases are influenced by the gold standards for the definition of infection and absence of infection in the tested populations and the spectrum of disease in the infected population (14). The standard for diagnosis of infection used in the present study was in the isolation of *M. paratuberculosis* from a fecal or tissue sample, a widely accepted standard. The standard for freedom of infection was the residence of the animal in a herd certified free of infection based on multiple independent (nonserologic) tests spanning several years. It is inappropriate to use test-negative cattle resident in *M. paratuberculosis*-infected herds for the estimation of diagnostic Sp because the chance of erroneously considering the animal noninfected is too great, i.e., such a "gold standard" is imperfect and inappropriate.

The spectrum of disease in the *M. paratuberculosis*-infected animals used in the present study was typical of that found in clinically normal adult Holstein cows raised in infected herds in

Wisconsin, the type of herds in which ELISAs are performed. Clinical samples from these same cattle have been used in the evaluation of other diagnostic tests for paratuberculosis (22–24). Se and Sp estimates found using the single manufacturer's S/P cutoff of 0.25 (45 and 97%, respectively) are similar to those reported by other investigators, suggesting that the spectrum of disease was at least comparable to that of animals typically used for paratuberculosis serologic test evaluations (11, 16, 25, 32). Arguably, these cases of paratuberculosis may not be typical of all truly infected cattle and do not include *M. paratuberculosis*-infected cattle that test negative by all available diagnostic methods; however, they were selected without bias for any single diagnostic test. Calculated LRs would be somewhat lower or higher if the spectrum of disease in the infected population was biased toward early- or very-late-stage infections, respectively.

Diagnostic laboratory medicine for animal agriculture is driven more by economics than is companion animal or human diagnostic laboratory medicine. Containment of testing costs and consideration of the economic consequences of the actions resulting from the diagnostic results are critical considerations in deciding which laboratory services to offer. These are factors that veterinary practitioners must consider in deciding what test to use in which circumstances.

The ELISA evaluated in this study produced quantitative results that were directly related to the likelihood that a dairy cow was infected with *M. paratuberculosis* (Table 1) and to the rate of positive results by using other diagnostic tests for paratuberculosis detection (Table 2). While the use of LRs in clinical epidemiology has long been advocated (18), few clinicians actually use them to estimate diagnostic probability (15). Additionally, as Feinstein nicely points out in his recent review, "the most important roles of technological tests today are in non-diagnostic clinical decisions," e.g., estimating prognosis (4). Acknowledging this, application of LR data has been simplified for practitioners by the creation of five categories of ELISA interpretation coupled to a recommended action scheme for paratuberculosis control in dairy herds (Table 3). These categories were derived somewhat arbitrarily but take into consideration the magnitude of the LRs and clinical experience with infected dairy herds. This scheme effectively

TABLE 3. Application of LR to the interpretation and use of ELISA results in *M. paratuberculosis*-infected dairy herds^{a,b}

S/P ratio ^c	Interpretation	Explanation and recommendation
0–0.09	Negative	Antibodies to <i>M. paratuberculosis</i> were not detected. The animal is either not infected or in a very early, undetectable stage of infection. Retesting in 6 to 12 months will increase confidence that the animal is free of paratuberculosis.
0.10–0.24	Suspect	Evidence of serum antibodies above normal background levels. The cow may be in the early stages of infection. Cows with this level of ELISA are roughly 5 to 15 times more likely to be <i>M. paratuberculosis</i> infected than the ELISA-negative cows. Cows with this result should give birth to their calves in a separate pen, and their colostrum should not be fed to calves.
0.25–0.39	Weak positive	Low level of serum antibodies to <i>M. paratuberculosis</i> but above the standard cutoff for a positive test. Odds are 16:1 or higher that this cow is infected. However, cows with weak-positive ELISA results may remain in good health for another lactation and are of limited infection transmission risk to the herd, provided colostrum, milk, and manure from these animals are kept away from calves.
0.40–0.99	Positive	Moderate level of serum antibodies to <i>M. paratuberculosis</i> . The odds that this cow is infected are at least 30:1. This cow is likely to be shedding the bacterium in its feces and milk and so should be culled from the herd soon and sold for slaughter only.
1.00–10.00	Strong positive	High level of serum antibodies to <i>M. paratuberculosis</i> . The odds that this cow is infected are over 200:1, and the animal is likely in the advanced stages of infection, shedding large numbers of the bacterium in feces and milk. It probably will soon develop clinical signs of the disease. The animal should be culled immediately and sold for slaughter only.

^a This interpretation scheme only is valid for herds of dairy cattle proven to be *M. paratuberculosis* infected by having at least one culture-confirmed case in an animal born and raised in the herd (herd prevalence above zero).

^b An earlier version of this table, based on preliminary data, appears on the Cornell website describing the New York State Cattle Health Assurance Program (<http://nyschap.vet.cornell.edu/>).

^c An ELISA S/P ratio of 0 indicates an antibody level equal to that of the negative control provided with the diagnostic kit. ELISA S/P ratio of 1 indicates an antibody level equal to that of the positive control provided with the diagnostic kit.

TABLE 4. Effect of pre-test probability and ELISA S/P value on the post-test probability^a of *M. paratuberculosis* infection

ELISA result	Sample midrange ELISA S/P value	Pre-test probability of <i>M. paratuberculosis</i> infection (estimated within-herd prevalence [%]) at:				
		1%	5%	10%	20%	25%
Negative	0.05	1	3	6	13	17
Suspect	0.20	9	35	53	72	77
Low positive	0.30	19	55	72	85	89
Positive	0.70	57	87	93	97	98
High positive	2.00	92	98	99	100	100

^a Pre-test odds \times LR = post-test odds, converted to probability, for example, for a cow with an ELISA S/P value of 0.70 in a herd with a 10% infection rate (probability that any randomly selected cow is infected): Step 1, convert pre-test probability (P) to pre-test odds: $[P/(1 - P)]$ or $1/(1 - 0.1) = 0.11$. Step 2, calculate the LR from the ELISA S/P value by the regression function reported in this paper: $LR = 265 \times 0.7^{2.03} = 128:1$. Step #3, multiply the pre-test odds with the LR to find the post-test odds: $0.11 \times 128 = 14.1$. Step #4, convert odds to probability [odds/(odds + 1)], or $14.1/15.1 = 93\%$.

lowers the cutoff for the identification of high-risk cattle to an S/P value of ≥ 0.10 (the manufacturer's S/P cutoff for a positive test is 0.25) and couples this with recommendations for low-cost interventions involving cows with such results to limit the spread of infection (segregated calving pens and discarding of colostrum). It also functionally raises the cutoff for cattle that are recommended to be culled from the herd to an S/P value of ≥ 1.00 , whereas the ELISA Sp is $\geq 99.9\%$ (Table 1), thereby limiting the rate of false-positive results and thus the economic impact on the herd owner caused by the mistaken sale and replacement of a noninfected cow. In this way, veterinary practitioners and dairy producers have a simple scheme for ELISA use in herds that is based on EBM and LRs without need of calculations or nomograms. Whether such a system will effectively work to control paratuberculosis in dairy herds is the subject of on going studies.

Practitioners may wish to incorporate pre-test probabilities of infection (the estimated within-herd paratuberculosis prevalence) together with the magnitude of the *M. paratuberculosis* ELISA result, translated into an LR, for a more precise estimation of the post-test probability of *M. paratuberculosis* infection. Table 4 illustrates the interaction of pre-test probability and ELISA S/P value on the post-test probability of *M. paratuberculosis* infection diagnosis. Other serologic tests for veterinary use could and should be interpreted based on LRs.

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Letter to the Editor

Calculation Method for Likelihood Ratios Dictates Interpretation

Recently, likelihood ratios were published for use as an aid in the clinical interpretation of test results obtained from a commercially available enzyme-linked immunosorbent assay (ELISA) for the detection of antibodies against *Mycobacterium avium* subsp. *paratuberculosis* in cattle (2). As described by Dr. M. T. Collins, likelihood ratios can provide diagnostic information for multiple levels of a diagnostic test result when results are reported on a continuous scale rather than as a dichotomous (positive and negative) outcome. Unfortunately, the author did not take advantage of this characteristic of likelihood ratios.

Because the formula used to calculate likelihood ratios presented in this article incorporates sensitivity and specificity estimates for the ELISA at several arbitrary cutoff points, the resulting comparisons remain dichotomous in nature. Further, the interpretation of these likelihood ratios as calculated is incorrect. For example, the likelihood ratio calculated at the 0.10 level indicates that cows with S/P ratios of ≥ 0.10 are five times more likely to be infected than noninfected. With this approach, all cows with ELISA S/P ratios of ≥ 0.10 are included in the comparison. Multiple likelihood ratios that allow for comparisons between different levels of S/P ratios were not calculated. Because the ELISA S/P ratios were not stratified into multiple categories, the likelihood ratios presented in Table 1 of Collins's article (2) are overestimates resulting from heterogeneity within the defined groups of interest.

Table 1 provides likelihood ratios based on the multiple-level approach proposed by Sackett et al. (3) and calculated from the data presented by Dr. Collins in his Table 1 (2).

The likelihood ratios calculated in this manner differ considerably from those presented by Dr. Collins, and differences in their subsequent interpretations may have significant clinical and economic consequences. For example, cows with ELISA S/P ratios of 0.10 to < 0.25 are 0.5 times as likely (i.e., half as likely) to be infected than noninfected, not "5 to 15 times" as likely to be infected, as Dr. Collins indicated in his Table 3 (2). Likelihood ratios demonstrating less-than-eightfold differences between comparison groups have been suggested to provide weak statistical evidence that a test result is indicative of a defined outcome (1).

Additionally, the exclusion of ELISA S/P ratios of < 0.00 from this analysis is a matter for concern. ELISA results from 1,097 animals, over one-third of the available results, were not considered. As this report indicates, ELISA S/P ratios of < 0.00 may constitute a significant proportion of the results obtained during herd screening. These results could have been incorporated into the analysis had the multiple-level approach for likelihood ratio calculations been utilized.

Likelihood ratios can provide additional diagnostic information for use in the clinical interpretation of the ELISA for *M. avium* subsp. *paratuberculosis* in cattle. However, care must be taken in order to interpret these values correctly and to make correct management recommendations based on them. These data do not support the interpretation of ELISA S/P ratios presented by Dr. Collins.

REFERENCES

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Alecia Larew Naugle
William P. Shulaw
William J. A. Saville
Thomas E. Wittum
Department of Veterinary Preventive Medicine
College of Veterinary Medicine
The Ohio State University
Columbus, Ohio

Beverly Byrum
Animal Disease Diagnostic Laboratory
Ohio Department of Agriculture
Reynoldsburg, Ohio

Author's Reply

Naugle et al. are correct in stating that Table 1 of the original paper (1) presents likelihood ratios (LRs) based on dichotomous interpretations of the ELISA results at multiple cutoffs and that this method of interpretation is different from a true multilevel LR analysis. That this method was used is indicated in the table by the term "ELISA S/P cutoff value" in the heading of the first column and by the footnote for "LR" in the last column, which gives the LR formula based on sensitivity and specificity.

LRs based on the same data used for the original article, including ELISA S/P results of < 0.00 , and calculated by using a multilevel approach are shown in Table 1 presented here.

The small number of infected cattle in the midrange categories limits the precision of the estimated LR. The negative association of ELISA S/P values with *Mycobacterium paratuberculosis* infection status in the 0.10 to 0.24 range is a function of as-yet-unexplained differences between U.S. and Dutch cattle.

To improve the precision of LR estimates, in the original analysis (1) I tried to include cattle from the full spectrum of infection (those shedding *M. paratuberculosis* in feces and

TABLE 1. Calculation of likelihood ratios using a multilevel approach

Range of ELISA S/P ratios	No. of infected cows	No. of noninfected cows	Likelihood ratio ^a
0.00– < 0.10	39	1,359	0.5
0.10– < 0.25	14	454	0.5
0.25– < 0.40	8	45	2.9
0.40– < 1.00	14	43	5.3
≥ 1.00	41	3	224.3
Total	116	1,904	

^a Indicates odds that a cow with an ELISA S/P ratio in a specific range will be infected, calculated as [(number of ELISA S/P ratios from infected cows in stratum)/(total number of ELISA S/P ratios from infected cows)]/[(number of ELISA S/P ratios from noninfected cows in stratum)/(total number of ELISA S/P ratios from noninfected cows)].

TABLE 1. LRs for for accurate diagnosis of *M. paratuberculosis* in cattle from ELISA S/P ratios^a

ELISA S/P ratio range	No. (%) of cows tested for <i>M. paratuberculosis</i> that were:		LR
	Infected	Noninfected	
<0.10 ^b	66 (46.15)	2,429 (81.67)	0.57
0.10–0.24	14 (9.79)	454 (15.27)	0.64
0.25–0.39	8 (5.59)	45 (1.51)	3.07
0.40–0.99	14 (9.79)	43 (1.45)	6.75
≥1.00	41 (28.67)	3 (0.10)	286.70
Total	143	2,974	

^a Calculation of the LRs is based on multilevel approach and uses the same data as the original study, including ELISA S/P results of <0.00.

^b The value <0.10 includes results for all animals with S/P values of <0.10, not just those within the limits of the 0.00 to 0.10 range, thereby accounting for the data perceived by Naugle et al. as excluded.

those not shedding the bacterium) and as large a population of controls from paratuberculosis-free herds as possible. This necessitated the use of sera from Dutch cattle, since The Netherlands had more readily available certified paratuberculosis-free herds than did the United States. The choice of cases and controls significantly affected the five-level LRs, as shown in Table 2.

These results argue for geographically specific LRs. Paratuberculosis ELISA specificity differences were also reported by Van Maanen et al. (Abstr. 7th Int. Colloq. Paratuberculosis, abstr. 139, 2002). More data are required to evaluate the necessity of geographically specific LRs.

I elected to use an S/P ratio of 0.25 as the cutoff because it is the one recommended by the kit manufacturer to define positive results and because I felt it was more pragmatic to use this cutoff than to alter it. Instead, I tried to divide the S/P values below the cutoff into two categories and those above the cutoff into three categories. With U.S. cattle in certified paratuberculosis-free herds used as the comparison control group, the two right-hand columns in Table 2, which includes five levels of LRs, show that dairy cattle in the S/P range of 0.10 to 0.24 are over five times more likely to be infected with *M. paratuberculosis*. This justifies the proposed low-cost intervention strategies to control the potential spread of paratuberculosis from animals such as those presented in Table 3 of the original paper (1). Based on the same definitions for cases and controls and Blume's criteria for strength of evidence based on LRs, dairy cattle with ELISA results in the 0.40 to 0.99 range show strong evidence of *M. paratuberculosis* infection.

While the statistical concerns raised by Naugle et al. are valid, the relationship between ELISA S/P ratios and the likelihood of *M. paratuberculosis* infection in dairy cattle (and the correlation of ELISA S/P positive results with those of other tests for *M. paratuberculosis* infection, shown in Table 2 of the original paper [1]) remains apparent, regardless of choice of

TABLE 2. Effects of different case-control combinations on the accuracy of LRs for cattle diagnosed as positive for *M. paratuberculosis* by ELISA S/P ratios

ELISA S/P ratio range	LR for:			
	Case A ^a -control A ^c	Case B ^b -control A	Case A-control B ^d	Case B-control B
<0.10	0.56	0.35	0.48	0.29
0.10–0.24	0.64	0.56	5.73	5.03
0.25–0.39	3.70	2.45	7.08	4.68
0.40–0.99	6.75	8.17	37.65	47.50
≥1.00	286.70	469.10	110.27	180.42

^a Case A, all cattle determined to be infected in original study ($n = 143$).

^b Case B, only fecal-culture-positive cattle ($n = 81$).

^c Control A, all noninfected cattle nonresident in the United States and The Netherlands ($n = 2,974$).

^d Control B, only noninfected U.S. cattle ($n = 760$).

cases and controls. The purpose of the study was to create a simple system for decision making by dairy producers and veterinary practitioners based on *M. paratuberculosis* ELISA S/P values that was founded on the principles of LR analysis. Feedback from end users about this system is favorable; it allows for the management of this infectious disease without excessive cost, particularly in comparison to the heretofore recommended "test-and-cull" programs for bovine paratuberculosis using tests with only positive and negative interpretations.

Modeling economic decision analysis will require precise, multilevel LRs based on a larger number of well-characterized cases of bovine paratuberculosis and appropriate controls, and it may be necessary to define them for specific geographic regions. One must also keep in mind the effects of biological variation in host response to exposure or infection with mycobacteria as well as technical variations in assay performance and not become too enamored with statistical precision when describing diagnostic test outcomes.

I thank my Ohio colleagues for pointing out the important difference between the dichotomous LRs I provided in Table 1 of the original paper (1) and the five-level LRs calculated according to the method of Sackett et al. Our exchange of ideas and the resulting expanded data analysis have shed more light on the calculation and use of LRs in bovine paratuberculosis ELISA interpretation.

REFERENCE

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M. T. Collins
 School of Veterinary Medicine
 University of Wisconsin—Madison
 2015 Linden Dr.
 Madison, WI 53706-1102