



Short communication

Herd-level prevalence of *Mycobacterium avium* subsp. *paratuberculosis* infection in United States dairy herds in 2007J.E. Lombard^{a,*}, I.A. Gardner^b, S.R. Jafarzadeh^c, C.P. Fossler^a, B. Harris^d, R.T. Capsel^d, B.A. Wagner^a, W.O. Johnson^e^a USDA, Animal and Plant Health Inspection Service, Veterinary Services, Centers for Epidemiology and Animal Health, Fort Collins, CO 80526-8117, USA^b Department of Health Management, Atlantic Veterinary College, University of Prince Edward Island, Prince Edward Island C1A 4P3, Canada^c Department of Medicine and Epidemiology, University of California, Davis, CA 95616, USA^d USDA, Animal and Plant Health Inspection Service, Veterinary Services, National Veterinary Services Laboratories, Ames, IA 50010, USA^e Department of Statistics, University of California, Irvine, CA 92697, USA

ARTICLE INFO

Article history:

Received 18 April 2012

Received in revised form 9 August 2012

Accepted 12 August 2012

Keywords:

Mycobacterium avium subsp.*paratuberculosis*

True herd-level prevalence

NAHMS

Composite fecal samples

ABSTRACT

Testing of composite fecal (environmental) samples from high traffic areas in dairy herds has been shown to be a cost-effective and sensitive method for classification of herd status for *Mycobacterium avium* subsp. *paratuberculosis* (MAP). In the National Animal Health Monitoring System's (NAHMS) Dairy 2007 study, the apparent herd-level prevalence of MAP was 70.4% (369/524 had ≥ 1 culture-positive composite fecal samples out of 6 tested). Based on these data, the true herd-level prevalence (HP) of MAP infection was estimated using Bayesian methods adjusting for the herd sensitivity (HSe) and herd specificity (HSp) of the test method. The Bayesian prior for HSe of composite fecal cultures was based on data from the NAHMS Dairy 2002 study and the prior for HSp was based on expert opinion. The posterior median HP (base model) was 91.1% (95% probability interval, 81.6 to 99.3%) and estimates were most sensitive to the prior for HSe. The HP was higher than estimated from the NAHMS Dairy 1996 and 2002 studies but estimates are not directly comparable with those of prior NAHMS studies because of the different testing methods and criteria used for herd classification.

Published by Elsevier B.V.

1. Introduction

Estimates of the true herd-level prevalence (HP) and within-herd prevalence of infectious agents provide baseline data for assessment of the progress of disease control programs. HP can be estimated using a decision rule based on results of samples (e.g. serum, feces, milk or tissues) from multiple individual animals in each herd or using single or multiple composite samples (e.g. bulk tank milk, milk filters or composite fecal samples) (NAHMS, 1997; Christensen and Gardner, 2000; Adaska and Anderson, 2003; Warnick et al., 2003; Van Kessel et al., 2011).

Ideally, test-based prevalence estimates should be adjusted for the sensitivity and specificity of the selected diagnostic method to enable valid comparisons of estimates among studies that use different tests.

There are few estimates of the HP of *Mycobacterium avium* subsp. *paratuberculosis* (MAP) in U.S. dairies. The National Animal Health Monitoring System's (NAHMS) Dairy 1996 study, which is the only national estimate of MAP prevalence, reported 21.6% of dairy operations infected (NAHMS, 1997). The study was designed to have a 90% confidence of detecting ≥ 1 positive cow in herds with $\geq 10\%$ of cows infected. Recent studies suggest that many herds have $< 10\%$ of cows infected and that the sensitivity of the serum ELISA is lower than what was believed in 1996. Hence, the reported prevalence estimate was conservative (Adaska and Anderson, 2003; Hirst et al., 2004; USDA,

* Corresponding author. Tel.: +1 970 494 7245; fax: +1 970 494 7228.
E-mail address: jason.e.lombard@aphis.usda.gov (J.E. Lombard).

2005). The NAHMS Dairy 2002 study also evaluated MAP infection but the low number of herds sampled and the fact that these herds were not randomly selected did not allow calculation of a national estimate (USDA, 2005). Of the 98 herds that were sampled in the 2002 study, 69 (70.4%) had at least one composite fecal sample that cultured positive for MAP (Lombard et al., 2006b), providing evidence that MAP prevalence in dairy operations was higher than previously reported.

Multiple studies have evaluated the use of composite fecal samples – samples from high traffic areas where manure from a large number of cows is deposited – to estimate herd-level and/or within-herd MAP prevalence in dairy herds (Raizman et al., 2004; Berghaus et al., 2006; Lombard et al., 2006b; Pillars et al., 2009; Aly et al., 2009). Although the number of composite fecal samples collected was not the same in all studies (2–6 samples/operation), all studies reported high herd sensitivity (HSe) (>70%) or detection of more than 70% of tested herds. False-negative results most likely occur when the within-herd prevalence and environmental load of MAP are low or when the only cows shedding MAP are non-lactating and not contributing feces to high traffic areas. False-positive culture results are considered rare and usually occur due to laboratory contamination.

The objective of the present study was to estimate the true HP of MAP infection in U.S. dairy herds in 2007 based on Bayesian analysis of culture results of composite fecal samples.

2. Materials and methods

2.1. Study overview

Serologic and culture data from the NAHMS Dairy 2002 study were used to construct a Bayesian prior for HSe of composite fecal cultures. The Bayesian prior for herd specificity (HSp) was based on expert opinion and allowed for rare false-positive results. The priors were then used to estimate true herd-level prevalence and predictive values of MAP from composite fecal samples collected from a nationally representative random sample of dairy operations during the NAHMS Dairy 2007 study.

2.2. Data sources

2.2.1. NAHMS Dairy 2007 study

In the 2007 study, a random sample of 3554 dairy operations from 17 major dairy states (California, Idaho, Indiana, Iowa, Kentucky, Michigan, Minnesota, Missouri, New Mexico, New York, Ohio, Pennsylvania, Texas, Virginia, Vermont, Washington, and Wisconsin) was eligible for participation. Of those 3554, 3304 (93%) were contacted by the National Agricultural Statistics Service. There were 2519 operations that completed the initial questionnaire and 1077 were eligible (30 or more cows on January 1, 2007) and consented to contact by a veterinary medical officer and potentially continue with the next phase of the survey. Of the 1077 operations that consented, 582 operations continued in the study and were eligible for testing of composite fecal samples for MAP (USDA, 2008a). Of the 582

eligible operations, 524 (90%) participated in the collection of composite fecal sample from areas on the dairy where manure from a majority of cows accumulated. Federal and state animal health officials collected samples from 6 different locations on each operation. Instructions were given to sample from areas specifically listed on the collection form (common alleyway, common pen, exit way from parlor, floor of holding pen, flush water, gutter cleaner, lagoon, manure pit, and manure spreader) rather than create samples from areas designated as 'other'. Samples were sent overnight on ice to the USDA-APHIS-VS, National Veterinary Services Laboratories (NVSL) in Ames, IA for culture. Samples were cultured on Herrold's egg yolk (HEY) agar (Becton Dickinson Diagnostic Systems, Sparks, MD). Two flasks containing HEY agar with Mycobactin J, two tubes of HEY agar with Mycobactin J, and one tube of HEY agar without Mycobactin J were inoculated. IS900 PCR was used to confirm positive cultures as MAP.

2.2.2. NAHMS Dairy 2002 study

Information from this study has been previously published (USDA, 2003; Lombard et al., 2006a,b). Briefly, a subset of 98 herds from 21 major dairy states (California, Colorado, Idaho, Illinois, Indiana, Iowa, Florida, Kentucky, Michigan, Minnesota, Missouri, New Mexico, New York, Ohio, Pennsylvania, Tennessee, Texas, Virginia, Vermont, Washington, and Wisconsin) was purposely selected from 1013 randomly selected operations that participated in the study. To be eligible to participate, operations had to have completed the initial questionnaire and have ≥ 30 dairy cows on January 1, 2002. Five composite fecal samples (compared with 6 in 2007) were collected by federal and state animal health officials from areas on these operations where manure accumulated from multiple adult cows. Samples were shipped on ice to the NVSL for culture. In contrast to the 2007 study, each sample was cultured by 3 methods (Herrold's egg yolk agar, Becton Dickinson Diagnostic Systems, Sparks, MD; BACTEC™ 460TB System, Becton Dickinson Diagnostic Systems; and ESP® Culture System II, Trek Diagnostic Systems, Cleveland, OH). If a sample was positive by any method, the sample was deemed positive. If one or more cow samples were culture positive for a particular operation, the herd was designated as culture-positive for MAP. The HSp of culture of composite samples was assumed to be perfect. Two of the 98 herds were excluded from estimation of the HSe of testing of composite fecal samples because of improper sampling. In one herd, samples were collected only from heifer areas and in the other, individual cow samples were collected rather than composite fecal samples. In order to estimate HSe based on less than 5 composite fecal samples, models created in SUDAAN® software (Release 10.0.1 2010, Research Triangle Institute, Research Triangle Park, NC) using the hypergeometric distribution were used to estimate HSe for 1–4 samples drawn without replacement from the population of composite fecal samples.

Serum ELISA testing was performed on the whole herd or a sample of the herd from these 96 operations. Blood samples were shipped overnight on ice to NVSL for testing. A commercially available ELISA kit (Paracheck, Biocor Animal Health, Omaha, NE) was used for testing as directed

Table 1

Herd-level sensitivity for composite fecal sampling to classify a herd as MAP-positive based on the number of composite samples tested and estimated within-herd true prevalence for operations ($n = 96$) in the NAHMS Dairy 2002 study.

No. of composite fecal samples	Herd-level sensitivity (%)			
	Low within-herd prevalence (>0–12.0%)	Moderate within-herd prevalence (>12.0–25.0%)	High within-herd prevalence (>25.0%)	All herds
5	55.6	87.0	89.7	77.2
4	55.6	87.0	82.8	74.7
3	48.1	78.3	75.9	67.1
2	44.4	65.2	69.0	59.5
1	33.3	56.5	58.6	49.4

by the manufacturer with the exception that samples were only tested in a single well, instead of in duplicate. Results were reported as negative or positive at the manufacturer-recommended threshold based on a computed ELISA score of 1.0. None of the herds used MAP vaccines although it is possible that some purchased cows might have been vaccinated. The apparent within-herd seroprevalence was calculated as the number of cows testing positive divided by the total number of cows tested for each herd. Based on the reported ELISA sensitivity of 29% and specificity of 99.7% (Collins et al., 2006), true within-herd prevalence for each herd was calculated from apparent prevalence (Rogan and Gladen, 1978). True prevalence was used to categorize herds in 3 groups based on the within-herd MAP estimate. True MAP within-herd prevalence ranges for the low, moderate, and high categories were >0–12%, >12–25%, and >25%, respectively.

2.3. Bayesian analysis

2.3.1. Prior distributions for HSe and HSp of composite fecal samples, and true HP

The priors for HSe were based on the conservative assumption that collection of 6 samples in 2007 would have the same HSe as 5 samples did in the 2002 study (Table 1) assuming a non-informative (beta 1,1) prior before the study was done. Therefore, the corresponding HSe prior was beta (62,19). The prior for HSp was based on the expert opinion of one of the coauthors (BH) who worked in the testing laboratory involved in the 2002 and 2007 studies. The HSp prior (beta (9999,1)) allowed for rare false-positive results (approximately 1 in 10,000) if the 6 samples were all truly negative (i.e., from a non-infected herd) and cross-contamination of 1 or more samples occurred in the testing laboratory. The HP prior was non-informative (beta (1,1)).

2.3.2. Bayesian model for herd prevalence

The model was based on the animal-level prevalence model of Branscum et al. (2004) where HP, HSe and HSp were substituted for their individual counterparts. Each herd's infection status was assumed to be independently Bernoulli distributed such that

$$Y_i \sim \text{Bernoulli}(p_i),$$

where Y_i was the test result ($Y_i = 1$ if positive and $Y_i = 0$ if negative) for herd i , and p_i was the probability of a positive result for herd i . Predictive values of herd-level

negative and positive test results were calculated because a population-based design was used.

The default priors for HP, HSe, and HSp were as described in Section 2.3.1. Posterior distributions were approximated using Markov-chain Monte Carlo (MCMC) methods in WinBUGS (Lunn et al., 2000). Posterior medians and 95% probability intervals (PI) were used for inferences and were based on 50,000 iterates after a burn-in of 5000 iterates. Convergence of the MCMC chain after the burn-in period was assessed by evaluation of trace plots and running multiple chains from different starting values.

Sensitivity analysis was done by evaluating the effects of changing priors for HSe and HSp. For HSe, an expert (R. Whitlock, personal communication) indicated that he might expect a slight increase (2–4%) in HSe in culturing 6 rather than 5 composite samples without considering the data in Table 1. This slight increase in HSe is reasonable given the trend in HSe evident in the 2002 data. The increase in HSe corresponded to the detection of 1 additional infected herd in each of the 3 prevalence categories and hence, the corresponding prior for HSe was beta (65,16). A second less optimistic prior for HSe (beta (56,25)) was used. The median HSe value for this prior was 69.3%, which approximated the lower limit of values reported in prior published studies (see Section 1). The effect of a pessimistic prior for HSp (beta (99,1)), which allowed for approximately 100 times the number of false-positives results compared with the default prior, was also evaluated.

3. Results

3.1. Test prevalence in 2007

Of the 524 tested herds in 2007, 369 (70.4%) had at least 1 of 6 samples that was culture positive for MAP (USDA, 2008b). In the culture-positive herds, the frequency distribution of 1–6 positive samples was 37 (10.0%), 38 (10.3%), 23 (6.2%), 48 (13.0%), 65 (17.6%), and 158 (42.8%), respectively.

3.2. Calculation of HSe using 2002 data

HSe was estimated using 1–5 composite fecal samples for 3 levels of herd prevalence and the results are presented in Table 1. As expected, as the number of composite samples tested increased and the estimates of within-herd true

Table 2

Posterior median and 95% probability intervals (PI) for herd-level prevalence and predictive values based on 3 different priors for herd sensitivity, a non-informative prior for herd-level prevalence, and a highly informative (beta 9999,1) prior for herd specificity.

Herd-level sensitivity (HSe) prior	Prevalence (95% PI)	Negative predictive value (95% PI)	Positive predictive value (95% PI)
Base model Beta (62,19): median HSe = 0.768	0.911 (0.815–0.993)	0.302 (0.025–0.593)	1 (0.999–1)
Lower sensitivity Beta (56,25): median HSe = 0.693	0.957 (0.866–0.998)	0.143 (0.006–0.427)	1 (1–1)
Higher sensitivity Beta (65,16): median HSe = 0.805	0.873 (0.785–0.978)	0.428 (0.079–0.687)	1 (0.999–1)

prevalence increased, the HSe increased. Culture of 5 composite fecal samples across all herds had a HSe of 77.2%.

3.3. Herd prevalence and predictive values

For the base model using a beta (62,19) prior for HSe, the posterior median herd prevalence was 91.1% (95% PI, 81.6% to 99.3%). Sensitivity analysis indicated that posterior estimates changed most with changes in the HSe prior (Table 2) and minimally with a decrease in the HSp prior (results not shown). For all models, the herd-level positive predictive values were close to 1 but median herd-level negative predictive values ranged from 13% to 43% depending on the HSe prior. Probability intervals for herd-level negative predictive value were wide.

4. Discussion

In the present study, we adapted code developed by Branscum et al. (2004) to allow estimation of true HP of *M. avium* subsp. *paratuberculosis* and the corresponding herd-level positive and negative predictive values in United States dairy herds, as a function of HSe and HSp of culture of 6 composite environmental samples.

Within-herd prevalence and the level of MAP contamination of the herd environment (environmental load) are important covariates affecting the HSe of composite fecal samples for herd level detection. Prevalence and environmental load may not be related depending on the association between shedding level in individuals and the proportion of heavy shedders. The environmental load and probability of a pen testing positive from composite fecal sampling was positively but not statistically correlated with the number of animals in the pen shedding in 3 low-prevalence herds (Smith et al., 2011). This relationship is further complicated by the phenomenon of MAP super-shedding (Whitlock et al., 2005; Aly et al., 2012).

Composite fecal sampling has received quite a bit of attention over the past few years. It has been used in a number of studies to evaluate MAP infection, primarily in dairy herds. This sampling method has also been used to evaluate *Salmonella* at the herd level for dairy operations and performed similarly to individual animal sampling (Lombard et al., in press). Composite fecal sampling is less costly and resource intensive than sampling individual cows to determine the herd MAP infection status while maintaining a relatively high sensitivity (~70%) (Lombard et al., 2006b).

The NAHMS Dairy 2002 study, to the authors' knowledge, collected composite fecal samples from the largest number of operations ($n = 96$) prior to the Dairy 2007 study and also had individual animal sampling to determine within-herd prevalence. The estimates of HSe for composite fecal sampling used in this study were modeled using the 2002 data. Pillars et al. (2009) used a HSe of 81% and HSp of 100% but only collected 2 composite fecal samples from each of the 94 sampled operations. This is likely the reason the reported true prevalence was only 49%; however, a value of 67% would have been obtained if a more appropriate HSe value of 59.5% (Table 1) based on the 2002 data was used in the analysis.

Smith et al. (2011) compared quarterly composite fecal sampling results from 3 low prevalence herds in the Northeastern U.S. to individual cow fecal culture results and reported a HSe of 40% when 6 composite sampled were collected. The NAHMS 2002 results suggest that collection of 5 composite fecal samples would result in a HSe of 55.6% – almost 50% higher than that calculated by Smith et al. (2011).

The posterior median HP (base model) was 91.0% (95% probability interval, 81.5 to 99.3%) and estimates were most sensitive to the HSe prior. The HP was higher than estimated from the NAHMS Dairy 1996 and 2002 studies but estimates are not directly comparable with those of prior NAHMS studies because of the different testing methods and criteria used for herd classification. Based on the NAHMS Dairy 2007 study and outcome of this modeling, it suggests the majority of dairy operations in the US are infected with MAP.

Conflict of interest statement

None.

Acknowledgements

Funding for data analysis was provided by the Johne's Disease Integrated Program (USDA-NIFA Coordinated Agricultural Project No. 2008-55620-18710). We thank Drs. Robert Whitlock and Michael Collins for providing expert opinion about sensitivity of composite sampling and the within-herd prevalence distribution, respectively. We also thank Patrick Camp, Philip Dykema, and Gabe Wilson for providing laboratory support in the culturing and identification procedures for the submitted samples.

Appendix A.

Code for the base model when herd sensitivity is modeled as independent of within-herd true prevalence. Abbreviations: hap=herd-level apparent prevalence; hse=herd sensitivity; hsp=herd specificity; hp=true herd-level prevalence; hppv=herd-level positive predictive value; and hnpv=herd-level negative predictive value.

```
model {
  y ~ dbin(hap, n)
  hap <- hp * hse + (1 - hp) * (1 - hsp)
  hse ~ dbeta(62, 19)
  hsp ~ dbeta(9999, 1)
  hp ~ dbeta(1, 1)
  hppv <- hse * hp / (hse * hp + (1 - hsp) * (1 - hp))
  hnpv <- hsp * (1 - hp) / ((1 - hse) * hp + hsp * (1 - hp))
}
```

References

- Adaska, J.M., Anderson, R.J., 2003. Seroprevalence of Johne's disease infection in dairy cattle in California, USA. *Prev. Vet. Med.* 60, 255–261.
- Aly, S.S., Anderson, R.J., Whitlock, R.H., Fyock, T.L., McAdams, S., Byrem, T.M., Jiang, J., Adaska, J.M., Gardner, I.A., 2012. Cost-effectiveness of diagnostic strategies to identify *Mycobacterium avium* subspecies *paratuberculosis* super-shedder cows in a large dairy herd using antibody enzyme-linked immunosorbent assays, quantitative real-time polymerase chain reaction, and bacterial culture. *J. Vet. Diagn. Invest.* 24, 821–832.
- Aly, S.S., Whitlock, R.H., Fyock, T.L., McAdams, S., Adaska, J.M., Jiang, J., 2009. Reliability of environmental sampling to quantify *Mycobacterium avium* subspecies *paratuberculosis* on California free-stall dairies. *J. Dairy Sci.* 92, 3634–3642.
- Berghaus, R.D., Farver, T.B., Anderson, R.J., Jaravata, C.C., Gardner, I.A., 2006. Environmental sampling for detection of *Mycobacterium avium* ssp. *paratuberculosis* on large California dairies. *J. Dairy Sci.* 89, 963–970.
- Branscum, A.J., Gardner, I.A., Johnson, W.O., 2004. Bayesian modeling of animal- and herd-level prevalences. *Prev. Vet. Med.* 66, 101–112.
- Christensen, J., Gardner, I.A., 2000. Herd-level interpretation of test results for epidemiologic studies of animal diseases. *Prev. Vet. Med.* 45, 83–106.
- Collins, M.T., Gardner, I.A., Garry, F.B., et al., 2006. Consensus recommendations on diagnostic testing for the detection of paratuberculosis in cattle in the United States. *J. Am. Vet. Med. Assoc.* 229, 1912–1919.
- Hirst, H.L., Garry, F.B., Morley, P.S., Salman, M.D., Dinsmore, R.P., Wagner, B.A., McSweeney, K.D., Goodell, G.M., 2004. Seroprevalence of *Mycobacterium avium* subsp. *paratuberculosis* infection among dairy cows in Colorado and herd-level risk factors for seropositivity. *J. Am. Vet. Med. Assoc.* 225, 97–101.
- Lombard, J., Beam, A., Nifong, E., Fossler, C., Koprak, C., Dargatz, D., Wagner, B., Erdman, M., Fedorka-Cray, P. Comparison of individual, pooled, and composite fecal sampling for detection of *Salmonella* on U.S. dairy operations. *J. Food Prot.*, in press.
- Lombard, J., Byrem, T., Wagner, B., McCluskey, B., 2006a. Comparison of milk and serum ELISA for diagnosis of *Mycobacterium avium* subspecies *paratuberculosis* infection in dairy cattle. *J. Vet. Diagn. Invest.* 18, 448–458.
- Lombard, J., Wagner, B., Smith, R., McCluskey, B., Harris, B., Payeur, J., Garry, F., Salman, M., 2006b. Evaluation of environmental sampling and culture to determine *Mycobacterium avium* subspecies *paratuberculosis* distribution and herd infection status on US dairy operations. *J. Dairy Sci.* 89, 4163–4171.
- Lunn, D.J., Thomas, A., Best, N., Spiegelhalter, D., 2000. WinBUGS – a Bayesian modeling framework: concepts, structure, and extensibility. *Stat. Comput.* 10, 325–337.
- NAHMS, 1997. Johne's Disease on U.S. Dairy Operations. USDA-APHIS-VS, CEAH, National Animal Health Monitoring System, Fort Collins, CO (#N245.1097).
- Pillars, R.B., Grooms, D.L., Woltanski, J.A., Blair, E., 2009. Prevalence of Michigan dairy herds infected with *Mycobacterium avium* subspecies *paratuberculosis* as determined by environmental sampling. *Prev. Vet. Med.* 89, 191–196.
- Raizman, E.A., Wells, S.J., Godden, S.M., Bey, R.F., Oakes, M.J., Bentley, D.C., Olsen, K.E., 2004. The distribution of *Mycobacterium avium* ssp. *paratuberculosis* in the environment surrounding Minnesota dairy farms. *J. Dairy Sci.* 87, 2959–2966.
- Rogan, W.J., Gladen, B.C., 1978. Estimating prevalence from results of a screening test. *Am. J. Epidemiol.* 107, 71–76.
- Smith, R.L., Schukken, Y.H., Pradhan, A.K., Smith, J.M., Whitlock, R.H., Van Kessel, J.S., Wolfgang, D.R., Grohn, Y.T., 2011. Environmental contamination with *Mycobacterium avium* subsp. *paratuberculosis* in endemically infected dairy herds. *Prev. Vet. Med.* 102, 1–9 (Epub 2011 July 19).
- USDA, 2003. Dairy 2002. Part III: Reference of Dairy Cattle Health and Health Management Practices in the United States, 2002. USDA-APHIS-VS, CEAH, National Animal Health Monitoring System, Fort Collins, CO (#N400.1203).
- USDA, 2005. Johne's Disease on U.S. Dairy Operations. USDA-APHIS-VS, CEAH, National Animal Health Monitoring System, Fort Collins, CO (#N427.0205).
- USDA, 2008a. Dairy 2007. Part III: Reference of Dairy Cattle Health and Management Practices in the United States, 2007. USDA-APHIS-VS, CEAH, National Animal Health Monitoring System, Fort Collins, CO (#N482.0908).
- USDA, 2008b. Johne's Disease on U.S. Dairies, 1991–2007. USDA-APHIS-VS, CEAH, National Animal Health Monitoring System, Fort Collins, CO (#N521.0408).
- Van Kessel, J.A., Karns, J.S., Lombard, J.E., Koprak, C.A., 2011. Prevalence of *Salmonella enterica*, *Listeria monocytogenes*, and *Escherichia coli* virulence factors in bulk tank milk and in-line filters from U.S. dairies. *J. Food Prot.* 74, 759–768.
- Warnick, L.D., Kaneene, J.B., Ruegg, P.L., Wells, S.J., Fossler, C., Halbert, L., Campbell, A., 2003. Evaluation of herd sampling for *Salmonella* isolation on midwest and northeast US dairy farms. *Prev. Vet. Med.* 60, 195–206.
- Whitlock, R.H., Sweeney, R.W., Fyock, T., Smith, J., 2005. MAP super-shedders: another factor in the control of Johne's disease. In: Manning, E.J.B., Nielsen, S.S. (Eds.), Proceedings of the 8th International Colloquium on Paratuberculosis. International Association for Paratuberculosis, Madison, WI, USA, p. 164.